Spatial competition and the dynamics of rarity in a temporally varying environment

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ABSTRACT

Questions: How does temporal variation in competitive advantage affect advance from rarity and species abundances in an individual-based ecology? In particular, how does the difference between the timescale of competitive invasion and the timescale of environmental periodicity interact with the spatial clustering underlying invasion to influence global population dynamics?

Features of model: We assume that two species compete preemptively for space in a two-dimensional environment. We categorize invasion of one of them as either the growth of one cluster of the rare species, or as growth of many clusters at the same time. Simulation of a constant environment identifies the characteristic timescale for a competitively superior species to invade and numerically dominate a resident species.

Manipulation of key variables: Given an endogenous timescale set by invasion in a constant environment, we introduced periodic temporal variation in competitive superiority by alternating the species’ propagation rates. We set the half-period of the environment much less than, roughly equal to, and much greater than the endogenous timescale. By manipulating habitat size and introduction rate, we simulated environments where successful invasion proceeds through growth of many spatial clusters, and where invasion can occur only as a single-cluster process.

Conclusions: In the multi-cluster invasion regime, rapid environmental variation produced spatial mixing of the species and non-equilibrium co-existence. The dynamics’ dominant response effectively averaged environmental fluctuation, so that each species could avoid competitive exclusion. Increasing the environment’s half-period to match the population-dynamic timescale let the (initially) more abundant resident repeatedly repel the invader. Periodic transition in propagation-rate advantage rarely interrupted the exclusion process when the more abundant species had a competitive advantage. However, at infrequent and randomly occurring times, the rare species could invade and reverse the density pattern by rapidly eroding the resident’s preemption of space.

In the single-cluster invasion regime, environmental variation that occurred faster than the population-dynamic timescale prohibited successful invasion; the first species to reach its...
stationary density (calculated for a constant environment) continued to repel the other during long simulations. When the endogenous and exogenous timescales matched, the species randomly reversed roles of resident and invader; the waiting times for reversal of abundances indicate stochastic resonance. For both invasion regimes, environmental fluctuation that occurred much slower than the endogenous dynamics produced symmetric limit cycles, alternations of the constant-environment pattern.

*Keywords*: ecological invasion, nucleation, population-dynamic timescale, spatial competition, stochastic resonance, temporal variation.

**INTRODUCTION**

Ecologists recognize that local dispersal induces spatially correlated population densities (Wilson, 1996, 1998; Ellner et al., 1998). These correlations can govern interaction frequencies and, consequently, impact the global dynamics of competing species (Tilman and Kareiva, 1997; Chesson, 2000; Dieckmann et al., 2000). Models for spatially detailed competition generally predict that conditions discriminating co-existence from exclusion can depend on the degree of local structure (e.g. Bolker and Pacala, 1999; Mágori et al., 2005; Caraco et al., 2006). More subtly, but no less significantly for our understanding of competition, interactions structured by local dispersal often increase the characteristic timescale of population dynamics well beyond that of mean-field models (Hurtt and Pacala, 1995; Lehman and Tilman, 1997; O’Malley et al., 2006a). For example, it may take much more time for an introduced, ecologically superior competitor to displace an inferior resident when individuals interact only at the neighbourhood scale, compared with a well-mixed dynamics (Gandhi et al., 1999; Korniss and Caraco, 2005; O’Malley et al., 2005).

Ecologists further recognize that environmental variation between generations can strongly influence competitors’ dynamics (Chesson, 1990; Ripa and Ives, 2003; Descamps-Julien and Gonzalez, 2005; Adler and Drake, 2008). Environmental fluctuations might reduce densities sufficiently that the chance of extinction increases. However, temporal variation, at some periodicities, might help prevent competitive exclusion (Hutchinson, 1961; Caswell and Cohen, 1995; Spencer et al., 2007; D’Odorico et al., 2008). In particular, asynchrony between competing species’ rates of growth in a fluctuating environment may promote co-existence via temporal niche differences (Chesson and Huntly, 1997; Snyder, 2007).

Only rarely have ecologists combined analyses of spatially heterogeneous populations with environmentally induced demographic fluctuation (Chesson, 1990; Holt and Barfield, 2003; Schoolmaster and Snyder, 2007). Our study takes an individual-based approach to this interaction; we model preemptive competition in an environment subject to temporal variation. Given an exogenous process that periodically reverses competitive superiority between two locally dispersing species, we ask how this individual-level variation affects the competitors’ global dynamics. The few previous individual-based ecological models with temporal variation (e.g. Holt et al., 2004) ordinarily consider a single population’s growth only. But recently Chan et al. (2009) studied competitive co-existence with a multi-type contact process subject to seasonal variation (see Discussion).

We categorize our results according to the difference between two fundamental timescales, in each of two competitive-invasion regimes. The first temporal scale is the expected time ($\tau$) required for a superior competitor to invade a resident species and advance to numerical dominance in a constant environment (endogenous timescale). The second temporal scale is the half-period of the environmentally induced demographic-rate
variation, $t_{1/2}$, the time that elapses between changes in competitive rank (exogenous timescale). Invasion regimes distinguish between a successful invader growing as a single spatial cluster and invasive growth distributed among many clusters (Korniss and Caraco, 2005). Similar analyses of interacting timescales in physical systems (e.g. Korniss et al., 2000; Buendia and Rikvold, 2008) suggest a set of metrics, and offer predictions paralleling our results on the ecology of spatial growth (see Discussion).

We organize the paper as follows. First, to compare the endogenous timescales of single-cluster and multi-cluster invasion in a constant environment, we briefly summarize an individual-based model for preemptive competition. Next, we introduce periodic variation in reproductive rates, and interpret a series of model simulations. To reveal the range of dynamic complexity emerging from the interaction of the timescale comparison with invasion regime, we employ some simple metrics that are functions of the difference between the competitors’ densities. Finally, we summarize the dynamics and place our work in context.

NUCLEATION AND ECOLOGICAL INVASION

Before addressing temporal variation, we recall our constant-environment model. It draws on nucleation theory for clustered growth. Nucleation theory is sometimes referred to as KJMA theory – for Kolmogorov (1937), Johnson and Mehl (1939), and Avrami (1940, 1941). Nucleation theory predicts timescales associated with spatial growth in multi-cluster systems. Recent applications of the theory span the physical (Rikvold et al., 1994; Ramos et al., 1999; Korniss et al., 2002) and biological sciences (Gandhi et al., 1999; Herrick et al., 2002; Jun et al., 2004; O’Malley et al., 2005; Zhang and Bechhoefer, 2006).

Suppose that a competitively inferior species advances to its self-regulated, equilibrium global density. Thereafter, individuals of a competitively superior species are introduced stochastically at a rate much smaller than the two species’ birth and death rates. Introduction occurs rarely but repeatedly. Given a sufficiently high initial density of the competitively inferior resident, and preemptive competition (Platt and Weis, 1985; Connolly and Muko, 2003; Yurkonis and Meiners, 2004; Rácz and Karsai, 2006), the resident can resist invasion for a long time before declining (perhaps suddenly) towards competitive exclusion. A dynamics where rare, random introduction events combine with strongly clustered growth of the invader – features of many plant and animal invasions (D’Antonio, 1993; Holway, 1998; Herben et al., 2000) – cannot be captured faithfully by either mean-field models (homogeneous mixing) or deterministic partial differential equations (Moro, 2001; Antonovics et al., 2006). However, nucleation theory provides a powerful framework to model growth of a locally dispersing invader.

Individual-based model of spatial competition

To model preemptive competition, we consider an $L \times L$ lattice with periodic boundaries. Each site represents the resources required to sustain a single individual. The local occupation number at site $x$ is $n_i(x) = 0, 1$ with $i = 1, 2$, representing the number of resident and invader individuals, respectively. An empty site may be occupied by species $i$ through propagation from a neighbouring site at rate $a_i n_i(x)$, where $a_i$ is the individual-level propagation rate for species $i$, and $n_i(x) = (1/\delta) \sum_{x' \in nn(x)} n_i(x')$ is the density of species $i$ in the neighbourhood about site $x$. $nn(x)$ is the set of the $\delta$ nearest neighbours of site $x$; in this study, we fix $\delta = 4$. Each species also may occupy an empty site through immigration from outside
the environment; each species has introduction rate $\beta$ per open site. Introduction and local propagation occur independently, so that species $i$ occupies an open site at total probabilistic rate $\beta_i + \alpha_i \eta_i$, for $i = 1, 2$. An occupied site opens through density-independent mortality of the individual; here each species has mortality rate $\mu$.

For constant environments, we take $\beta_i / H < \alpha_i < \alpha_2$; introduction is rare, and invaders have a per-individual reproductive advantage. In simulation, we track the time-dependent global density of each species, $\rho_i(t) = (1/L^2) \sum_x n_i(x, t)$; $i = 1, 2$. We define the resident’s lifetime $\tau$ as the first passage time of the resident’s global density to 1/2 its initial (hence, quasi-equilibrium) density $\rho_1^*$. For given parameters, we take the resident’s mean lifetime $\langle \tau \rangle$, invasion time for simplicity, as the characteristic timescale of the population dynamics. The 1/2 is arbitrary; we choose it since the invader’s global density, $\rho_2(t)$, should just exceed the resident’s density when $\rho_1(t) = \rho_1^*/2$. The term ‘invasion time’ fits the constant-environment dynamics, since the superior invader always becomes common (though the elapsed time varies randomly). Since we introduce individuals of each species, competitive exclusion does not imply true extinction. Rather, we equate exclusion with a small global density $O(\beta)$, where the common introduction rate satisfies $10^{-8} \leq \beta \leq 10^{-4}$.

As simulation proceeds, individual invaders occasionally appear interspersed among the initially common residents. An invader lacking access to empty, neighbouring sites may die before propagating. If a site opens in the local neighbourhood, the invader may colonize it. However, the resident’s greater local density may compensate for its lower individual-level propagation rate, so that the resident species has the greater chance of colonizing the empty site. Consequently, most small invader clusters shrink and disappear. Residents, though weaker competitors, can avert exclusion for extended periods, since preemptive competition constrains invader growth. Invaders can succeed only after they generate a cluster sufficiently large that it tends to grow at its periphery (Yasi et al., 2006; Allstadt et al., 2007). We say a nucleation event occurs when a cluster’s radius first reaches $R_{\text{crit}}$.

**Single-cluster and multi-cluster invasion**

For given parameters, let $D(\alpha_1, \alpha_2, \mu, \beta)$ represent the average distance separating invader clusters in an infinitely large (hypothetical) environment. If the actual environment is sufficiently small, $L \ll D$, then invasion almost always proceeds through the growth of a single cluster (SC) invasion. However, when $L \gg D$, invasion involves many invader clusters [multi-cluster (MC) invasion]. Furthermore, suppose that we fix linear habitat size $L$, as well as other parameters – except the introduction rate. There then exists a characteristic value of the introduction rate $\beta$ (now governing $D$) such that MC invasion crosses over to the SC pattern for levels of $\beta$ less than the characteristic value (O’Malley et al., 2006a).

We can distinguish SC and MC invasion, and identify ecological implications, within the framework of homogeneous nucleation (Korniss and Caraco, 2005). We outline the essentials in Appendix 1; here we specify how the endogenous timescale differs between modes of invasion.

For SC invasion, the lengthy waiting time until the first successful cluster nucleates dominates the lifetime $\tau$ of the competitively inferior resident. Differences between the competitors’ demographic rates govern the time required for the cluster to grow, but growth time is short compared with the exponentially distributed waiting time for nucleation (O’Malley et al., 2006a). Hence if $t_{\text{nuc}}$ is the random waiting time until nucleation occurs (the first
invader cluster with a radius as large as $R_{\text{crit}}$, then $\langle \tau \rangle = \langle t_n \rangle$. From Appendix 1, $\langle t_n \rangle = (L^2 \beta)^{-1}$. Therefore, the characteristic timescale, $\langle \tau \rangle$, varies inversely with both habitat area and the introduction rate (O’Malley et al., 2005). To compare invasion modes, we can say the endogenous timescale $\langle \tau \rangle$ scales with $\beta^{-1}$ in the SC regime. Single-cluster invasion predicts that an invader maintains a low mean density, perhaps below a detection threshold, for an uncertain length of time – which can be quite long. Then, stochastically, a cluster grows beyond critical radius, and the invader drives the resident towards competitive exclusion.

For large environments (or large introduction rates), $L \gg D$; the environment’s size exceeds the typical distance between invader clusters. Consequently, many randomly nucleated, expanding clusters drive the resident’s decline. In the limit of a large environment, we can approximate the time-dependent global densities by Avrami’s law (see Appendix 1), which provides a functional form for global densities during MC invasion. Note that for MC invasion, $\langle \tau \rangle$ is essentially independent of the habitat size (Korniss and Caraco, 2005). Finally, for yet larger $\beta$, invader clusters coalesce immediately, and nucleation theory breaks down; in fact, homogeneous mean-field equations may apply as the clusters mix (O’Malley et al., 2006b).

Invasion time $\langle \tau \rangle$, our population-dynamic timescale, depends on the introduction rate $\beta$, and this dependence differs between invasion regimes. For simulated MC invasion, the timescale of the invader’s advance to numerical superiority scales as $\langle \tau \rangle \sim \beta^{-0.30 \pm 0.015}$, close to $\beta^{-1/3}$, the value predicted by Avrami’s law (O’Malley et al., 2006a). However, for any finite environment with linear size $L$, a sufficiently small $\beta$, such that $D(\mu_1, \mu_2, \mu, \beta) \sim \beta^{-1/3} \gg L$, implies SC invasion. In the SC regime, the average lifetime (and its standard deviation) scales as $\langle \tau \rangle \sim \beta^{-1}$, reflecting the Poisson nature of an invader cluster’s nucleation (see Appendix 1). For an extended analysis of invasion time versus invasion regime, see O’Malley et al. (2005).

**TEMPORAL VARIATION AND SPATIAL COMPETITION**

Most communities experience some level of temporal variation (Descamps-Julien and Gonzalez, 2005; Greenman and Norman, 2007; Adler and Drake, 2008; Loreau and de Mazancourt, 2008), and these environmental fluctuations may impact ecological and evolutionary processes (Caswell and Cohen, 1995; Ives, 1995; Chesson, 2000; Neubert et al., 2000; Travis et al., 2005; Altizer et al., 2006; Spencer et al., 2007). Given our model’s behaviour in a constant environment, we ask how the dynamics responds to periodic temporal variation. The distinction between single- and multi-cluster invasion proves useful in a periodically varying environment (Korniss et al., 2002).

**Time-dependent propagation rates**

To study a symmetric demographic response to environmental fluctuation, we modified our individual-based model by setting $\alpha_i(t) = \alpha + \Delta(t)$, and $\alpha_f(t) = \alpha - \Delta(t)$. We report results where $\Delta(t)$ is a square-wave with amplitude $\Delta$ and half-period $t_{1/2}$ (see Buceta et al., 2004). Since each species has the same individual mortality rate (we fix $\mu = 0.1$ throughout this paper), the environment alternately favours first one then the other species via the propagation-rate difference. Here we let $\alpha = 0.75$ and $\Delta = 0.05$, so that $[\alpha_i(t), \alpha_f(t)]$ alternates between [0.8, 0.7] and [0.7, 0.8]. We selected the invasion regime, MC or SC, by choosing habitat size $L$ and introduction rate $\beta$ appropriately. Given $(\alpha, \mu)$, these parameter selections also determine the population-dynamic timescale $\langle \tau \rangle$. In simulation, we varied the exogenous timescale, the environmental half-period $t_{1/2}$, from near-zero to levels far exceeding $\langle \tau \rangle$. 


For the parameter values that we used, summed global densities \( \rho_{\text{tot}}(t) = \rho_1(t) + \rho_2(t) \) exhibit little temporal variation. However, the difference in the species’ densities, \( m(t) = \rho_1(t) - \rho_2(t) \), proves relevant and informative (see Ripa and Ives, 2003). \( m(t) \), with range \( m(t) \in [-1, 1] \), indicates numerical superiority at time \( t \). If \( m(t) > 0 \) persistently, species 1 dominates numerically; if \( m(t) < 0 \) persistently, species 2 dominates. If \( m(t) = 0 \), or fluctuates about 0, the competitors co-occur with roughly equal global densities. Figure 1 shows two time series from our simulations. These illustrate the role of the density metrics. Note that \( m(t) \) plays a role here analogous to the order parameter in ferromagnetic systems (Lo and Pelcovits, 1990; Tomé and de Oliveira, 1990; Chakrabarti and Acharya, 1999).

Since we are interested in the competitive system’s behaviour in a periodically changing environment, a natural quantity to monitor is the period-averaged density difference (Tomé and de Oliveira, 1990; Sides et al., 1998a, 1998b; Korniss et al., 2002; Buendia and Rikvold, 2008):

\[
Q = \frac{1}{2T_{1/2}} \int_a^b m(t) dt,
\]

where an environmental period begins at \( t = a \) and ends at \( b = a + 2T_{1/2} \). The period-averaged density difference \( Q \) is also referred to as the dynamic order parameter (Sides et al., 1998a; Korniss et al., 2000). For \( T_{1/2} \gg \tau \), the environment changes slowly; decay of the respective resident’s density is complete, and the system approaches competitive exclusion in each half-period. Consequently, \( m(t) \) reaches a limit cycle symmetric about zero, and in turn \( Q = 0 \). On the other hand, for \( T_{1/2} \lesssim \tau \), the two species’ densities do not have enough time to ‘switch’ during each environmental half-period. Therefore, \( m(t) \) reaches an asymmetric limit cycle, and in turn, \(|Q| > 0 \). However, as we shall see in the next subsection, the scenarios can be more subtle than these two naïve cases; the quantity \( Q \) provides an important measure of difference in abundance in periodically varying environments. In the simulations, we evaluated \( Q \) by averaging \( m(t) \) through each full period, and we define \( \langle Q \rangle \) as the average of \( Q \) over many successive periods. We also calculated the average absolute value, \( \langle |Q| \rangle \), from the same data. Finally, we estimated a scaled, among-period variance of \( Q \) (Sides et al., 1999; Korniss et al., 2000; Buendia and Rikvold, 2008):

\[
X_Q^2 = L^2 \left( \langle Q^2 \rangle_L - \langle |Q| \rangle _L^2 \right)
\]

for different habitat sizes \( L \).

These metrics do more than describe patterns in the \( \rho_i(t) \). They identify transitions (‘crossovers’) in the competitive dynamics as functions of \( T_{1/2} \). That is, \( \langle |Q| \rangle \) declines rapidly towards zero, and \( X_Q^2 \) peaks, at critical values of \( T_{1/2} \) where transitions in the dynamics occur (Sides et al., 1998a; Korniss et al., 2002; Robb et al., 2007; Buendia and Rikvold, 2008).

We also estimated the logarithmic growth rate of species \( i (i = 1, 2) \) as:

\[
G[\rho_i(t)] = \ln L^2 \rho_i(t + 1) - \ln L^2 \rho_i(t) = \ln [\rho_i(t + 1)/\rho_i(t)].
\]

We were particularly interested in growth rates at low density, where clusters of the rare species should be small.

**Multi-cluster invasion in a periodic environment**

To begin, we let competitive advantage alternate periodically with \( \beta = 10^{-4} \), and \( L = 90, 128, 180, \) and 256. In a constant environment, these combinations of introduction rate and
habitat size assure MC invasion with a lifetime $\langle \tau \rangle \approx 1240$. Note that for MC invasion, the endogenous timescale $\langle \tau \rangle$ is essentially independent of habitat size (Korniss and Caraco, 2005). We advance simulation time as Monte Carlo system steps. That is, during a single time unit we randomly select $L^2$ sites for updating via the individual-based model. Then each site, on average, updates once per time step.

Figure 2A plots $\langle |Q| \rangle$ as a function of $t_{1/2}$: $\langle |Q| \rangle$ declines towards zero for both very short ($t_{1/2} \ll \langle \tau \rangle$) and very long ($t_{1/2} \gg \langle \tau \rangle$) periods. At intermediate periodicity, when $t_{1/2}$ is ‘less than but within the same order of magnitude’ as $\langle \tau \rangle$, we write as $t_{1/2} \approx \langle \tau \rangle$, $\langle |Q| \rangle$ takes distinctly non-zero values. Hence the competitors’ dynamics should exhibit two major transitions as $t_{1/2}$ increases – implying three different dynamical behaviours. Figure 2B verifies the transitions. The scaled variance $\sigma^2$ has two peaks (more pronounced as habitat size increases), each indicating a crossover from one dynamics to another.
Consider $t_{1/2} \ll \langle \tau \rangle$ first. The rapidly oscillating environment interrupts the dynamics’ relaxation towards equilibrium each half-period (see Buceta et al., 2004). Hence, once both species have entered the habitat, neither can exclude the other (in ecological time), and we observe non-equilibrium co-existence. Figure 3A shows that the difference in global densities responds regularly, but with small magnitude, to each brief half-period’s competitive asymmetry. Among periods the dynamics behaves, in effect, as if the time-varying environment had been replaced by its average (in physicists’ terminology, the rapid, periodic environmental variation ‘anneals’ propagation rates) (see inset of Fig. 3A). In turn, the period-averaged density difference (or dynamic order parameter) $Q$ fluctuates randomly about zero (Fig. 3D). Stated very carefully, we may say that reproductive advantage alternates so rapidly that the competitive interaction approaches the indeterminate case.
Fig. 3. Multi-cluster regime: segments of the density difference \( m(t) \) as a function of time, and period-averaged density difference \( Q \) plotted for a large number of consecutive periods. \( \beta = 10^{-4} \) and habitat size \( L = 128 \). (A) \( t_{1/2} = 100 \ll \langle \tau \rangle \). Note scale of ordinate. Inset: Long-term plot, indicating non-equilibrium co-existence with roughly equal averaged densities. (B) \( t_{1/2} = 800 \approx \langle \tau \rangle \). Invasion–exclusion cycles. Inset: Long-term plot. (C) \( t_{1/2} = 4000 \gg \langle \tau \rangle \). The system reaches single-species equilibrium during each, lengthy half-period. (D) \( t_{1/2} = 100 \ll \langle \tau \rangle \); each period lasts 200 time units. (E) \( t_{1/2} = 800 \approx \langle \tau \rangle \); each period lasts 1600 time units. Note the large transitions in \( Q \) (\( m \) changes sign). (F) \( t_{1/2} = 4000 \gg \langle \tau \rangle \); note limited scale of the ordinate.
where the two species have identical demographic parameters. Over the long term, the 
species slowly exchange roles of more and less abundant \( m(t) \) changes sign, inset of Fig. 3A; see Fig. 10A), but each competitor’s global density exceeds the background level main-
tained by introduction.

The non-equilibrium co-existence induced by rapid environmental fluctuation inhibits large-scale spatial order. Species 2 is rare initially. Introduced individuals of species 2 gener-
ate relatively small, randomly located clusters. These clusters ordinarily expand and con-
tract a bit during consecutive half-periods, and some (by chance) grow large enough to 
persist at length. That is, some clusters grow sufficiently large to maintain their single-
species equilibrium within, and localize between-species competition to their periphery (Allstadt et al., 2007, 2009), promoting cluster longevity. Consequently, neither competitor pro-
duces the single large-scale cluster that precedes exclusion of the other species (Gandhi et al., 
1998, 1999; Yasi et al., 2006). Figure 4 provides detail of the spatial system at the completion of 
consecutive half-periods. During the first half-period, species 2 had the greater propagation 
rate and advanced (Fig. 4A), whereas species 1 had the advantage during the second half-
period (Fig. 4B). Local correlations are evident, but the system does not exhibit ordering at 
extended distances and co-existence results.

Second, we increase the half-period to match invasion time, so that \( t_{1/2} = \langle \tau \rangle \) for MC parameters. Compared with the rapid environmental fluctuation just described, the reduced 
frequency of competitive reversal allows the dynamics to respond more strongly to the exogenous signal each half-period. Within most (but not all) intervals of length \( 2t_{1/2} \), the density-difference cycles from \( m = 0.9 \) to \( m = 0 \), and then returns to \( m = 0.9 \) (Fig. 3B). That is, species 2 advances from rarity during the half-period when it had the greater individual-level propagation rate, so that at the end of the half-period global densities are
roughly equal (Fig. 5A). During the next half-period, the now superior species 1 grows and excludes species 2 (Fig. 5B). We refer to this pattern as an invasion–exclusion cycle. Since $\langle \tau \rangle$ is defined as the mean time elapsing until an inferior, resident species’ density is halved, the system’s behaviour when $t_{1/2}$ and $\langle \tau \rangle$ are approximately equal follows from our understanding of invasion in a constant environment. The resident species maintains a greater density over the course of an entire environmental period, although it has no greater average propagation rate. Indeed, for thousands of consecutive periods, one species advances during one ‘season’, only to be excluded the next – despite the two species’ identical period-averaged demographic rates.

However, rough equivalence of the competitors’ average densities appears over extended timescales. Although the same invasion–exclusion cycle repeats for a long time, noise can generate spontaneous density fluctuations large enough to ‘flip’ global densities so that the rare species becomes the more abundant one. Figure 3E shows that $Q$, at apparently random times, switches from $Q = 0.6$ to $Q = -0.6$, reversing roles of common and rare species in less than $10^2$ periods. This lengthy-timescale switching between the two phases, each of which has a single, numerically dominant species throughout, is the manifestation of spontaneous symmetry breaking in an interacting particle system (Binder and Heerman, 1997). This process, often referred to as a dynamic phase transition, has been well studied in physical systems with local interactions (e.g. Korniss et al., 2000; Machado et al., 2005; Robb et al., 2007). Applying an insight from these studies yields an interesting ecological prediction: as a function of habitat size, the switching time between the two symmetric phases should increase faster than any power law. That is, the waiting time for fluctuations spanning the habitat, and consequently capable of switching the species’ abundances, should increase exponentially with the size of the habitat (Goldenfeld, 1992). Figure 6 demonstrates this effect in our model. Matching of the

Fig. 5. MC regime, $t_{1/2} = 800 \approx \langle \tau \rangle$. Open: white; species 1: grey; species 2: black. (A) Lattice at end of half-period during which species 2 had propagation advantage; $m(t) = 0$. (B) Lattice at end of next half-period, when species 1 had propagation advantage; $m(t) = 0.9$ at completion of invasion–exclusion cycle.
fundamental timescales gives rise to symmetry breaking, i.e. long-term dominance of one of the species (in very large habitats) despite the same mean demographic rates. At any given time, it is highly probable that one species holds numerical superiority; the other is introduced and then excluded each period. But one cannot predict, \textit{a priori}, which species will be common.

Third, when the half-period exceeds the timescale of invasion, $t_{1/2} \gg \langle \tau \rangle$, the currently superior species excludes the other following each environmental change. Invader clusters grow, coalesce, and the invader’s global density advances to single-species equilibrium before roles reverse, i.e. competitive exclusion is completed each half period. The symmetric limit cycle of $m(t)$ and associated $Q$ values are plotted in Figs. 3C and 3F, respectively.

\textbf{Growth rates and global densities}

Our model’s preemptive competitors are equally subject to self-regulation and interspecific competition; hence we observe near constancy of summed densities during simulation (Fig. 1C and D). Given the symmetry in the temporally varying propagation rates, each growth rate $G(\rho_i)$, averaged over time, should be zero. But the difference between the dynamics of non-equilibrium co-existence and invasion–exclusion cycles (i.e. the absence/presence of time for the system to relax to equilibrium) might be revealed in growth-rate variation. Large samples of per capita growth rates reflect this difference only subtly. Figure 7 plots $10^3$ logarithmic growth rates ($G(\rho_i(t))$) of the initial resident (species 1) and
Fig. 7. Logarithmic growth rates, as functions of density, MC regime. Species 1 (plots A, C) is resident, with initial density ≈ 0.86. Species 2 (plots B, D) is the invader, with initial density 0. (A, B) \( t_{1/2} = 100 \ll \langle \tau \rangle \); (C, D) \( t_{1/2} = 800 \gg \langle \tau \rangle \). Each plot includes 1000 estimates sampled regularly at intervals of 500 time units. When the environment oscillates relatively rapidly (plots A, B), non-equilibrium co-existence allows each species to traverse the same range of global densities over \( 10^5 \) time units. The growth-rate plots are nearly identical. When timescales match (plots C, D), we observed the invader more often at low density, a consequence of initial conditions. Furthermore, matching timescales (\( t_{1/2} = 800 \)) indicate invasion–exclusion cycling in invader growth; growth-rate variability is relatively large for \( \rho_2(t) < 0.2 \).
initial invader (species 2), for both \( t_{1/2} \ll \langle \tau \rangle \) and \( t_{1/2} \approx \langle \tau \rangle \). We sampled each species regularly at intervals large enough to ensure statistical independence.

Figures 7A and 7B reiterate the point that under rapid environmental oscillation, non-equilibrium co-existence allows each species, over lengthy timescales, to explore the same range of global densities. A comparison of Figs. 7A and 7C shows that for species 1 (the initial resident) neither the spatially averaged growth rates nor their levels of variability depend on that species’ global density. For matching timescales (Figs. 7C and 7D), we observed the invader’s growth rate more often at low density, a result of its initial rarity. Figure 7D indicates invasion–exclusion cycles of invader growth; growth-rate variability is larger for lower levels of \( \rho_2(t) \). This pattern of variability is likely more than an effect of increased samples at low density. Invasion–exclusion cycles imply negative growth rates for species 2 even at low density when it has the lower propagation rate and faster growth when it is rare during half-periods when it has the higher propagation rate.

**Single-cluster invasion in a periodic environment**

Setting \( \beta = 10^{-6} \) and letting \( L = 64 \) or 128 ensures SC invasion in a constant environment. Waiting times for successful introduction and invader nucleation are much longer for SC than for MC invasion (O’Malley et al., 2005). For \( L = 64 \), we estimate invasion time \( \langle \tau \rangle = 14,000 \), and estimate \( \langle \tau \rangle = 6000 \) for \( L = 128 \). Therefore, we simulated a wider range of \( t_{1/2} \) for SC invasion than for the MC case.

Figure 8A plots \( \langle |Q| \rangle \) against the environment’s half-period. \( \langle |Q| \rangle \) takes distinctly non-zero values at small \( t_{1/2} \), and declines towards zero as period length increases. The scaled variance \( X_Q^L \) (Fig. 8B) has a single peak, for a given habitat size \( L \), where \( t_{1/2} = \langle \tau \rangle \). The two results concur; the dynamics exhibits a single major transition as \( t_{1/2} \) increases, crossing over where the exogenous and endogenous timescales match.

For a rapidly alternating environment \( (t_{1/2} \ll \langle \tau \rangle) \), the rare species never can invade successfully (over \( 10^5 \) to \( 10^6 \) periods). Invader appearance and minimal spatial growth occur when the rare species has the greater propagation rate (see Figs. 9A and 9D). But exclusion of the invader follows quickly after competitive advantage switches to the resident. The frequency of environmental change, combined with a very small introduction rate, inhibits formation of a critically sized cluster of the rare species. Given an initially common species, the likelihood that the invader can establish a persistent cluster before losing competitive advantage and being excluded remains quite small.

When the environment changes very slowly, so that \( t_{1/2} \gg \langle \tau \rangle \), sufficient persistence of its propagation-rate advantage lets the currently superior competitor invade and exclude the other species in almost every half-period. Species’ densities, hence \( m(t) \), exhibit limit cycles, and \( Q \) fluctuates about zero (see Figs. 9C and 9F).

For rapid environmental change, the invader cannot disrupt the resident’s spatial order, even though the species are ‘on average’ equivalent. For very slow environmental change, invasion and invader growth from rarity to numerical abundance are assured.

**Single-cluster invasion and stochastic resonance**

Finally, consider matching timescales \( (t_{1/2} \approx \langle \tau \rangle) \) in the SC regime. Figure 9B shows that \( m(t) \) responds to environmental periodicity but does not track it faithfully. In some half-periods the species with temporary competitive advantage advances from rarity and excludes
Fig. 8. Metrics for single-cluster invasion. (A) Average of absolute value of $Q$ as a function of environmental half-period $t_{1/2}$; $\beta = 10^{-6}$, habitat size $L = 64$, 128; $\langle \tau \rangle \approx 14,000$ and 6000, respectively. (B) The scaled variance of $Q$. Single peak indicates value of $t_{1/2}$ where transition in competitive dynamics occurs. Inset: Rescaling shows singular phenomenon.
Fig. 9. Single-cluster regime: segments of the density difference $m(t)$ as a function of time, and period-averaged density difference $Q$ plotted for $10^3$ consecutive periods. $\beta = 10^{-6}$; habitat size $L = 128$; $\langle \tau \rangle \approx 6000$. (A) $t_{1/2} = 400 \ll \langle \tau \rangle$. Note limited scale of the ordinate. Inset: Long-term plot, showing that species 2 never advanced significantly during simulation. (B) $t_{1/2} = 8000 \approx \langle \tau \rangle$. Stochastic resonance. (C) $t_{1/2} = 20,000 \gg \langle \tau \rangle$. The system reaches single-species equilibrium during each, lengthy half-period. (D) $t_{1/2} = 400 \ll \langle \tau \rangle$; each period lasts 800 time units. Note limited scale of the ordinate. (E) $t_{1/2} = 8000 \approx \langle \tau \rangle$; each period lasts 16,000 time units. (F) $t_{1/2} = 20,000 \gg \langle \tau \rangle$; compare with Fig. 3F.
the resident. But successful invasion in any given period is uncertain. That is, the dynamics unpredictably switches between exclusion of one species and exclusion of the other (see Fig. 9E). Hence the competitive system exhibits a transitional behaviour between (essentially) no invasion ($t_{1/2} \ll \langle \tau \rangle$) and almost certain invasion each half-period ($t_{1/2} \gg \langle \tau \rangle$).

Erratic switching between competitive-exclusion states when the two timescales match suggests stochastic resonance (Gammaitoni et al., 1995). Consider a non-linear system with two locally stable, equilibrium nodes. To this we add a relatively small periodic input (alternating propagation rates). Depending on the initial conditions, the dynamics will then oscillate gently around one or the other equilibrium; the system will not ‘switch’ between alternative equilibria.

Next we add noise. The system now has a non-zero switching probability. That is, with stochasticity added, the dynamics may move from one attractor to the other. This switching in the presence of noise, whether randomly erratic or approaching the pattern of the periodic input, defines stochastic resonance (Marchesoni et al., 1996; Huppert and Stone, 1998; Sides et al., 1998a; Vilar and Solé, 1998; Korniss et al., 2002).

Our competition model does not exhibit bistability. In a constant environment, exclusion of the inferior species is the sole positive equilibrium. However, the competitively inferior resident can resist invasion for a long time because competition is preemptive (Korniss and Caraco, 2005). The constant-environment model possesses a stable equilibrium and a metastable ‘quasi-equilibrium’. Over a half-period, the latter may prevent the system from moving to the stable equilibrium. This sort of asymmetric system also can exhibit stochastic resonance when stability properties periodically reverse (Stocks et al., 1993; Leung and Néda, 1999).

To demonstrate stochastic resonance in the simulation data, we estimated frequencies of waiting times that elapsed between consecutive occurrences of equal species’ abundance (i.e. consecutive times when $m(t) = 0$). Designate the random waiting time $t_r$. For square-wave external variation, Korniss et al. (2002) assume that the time for growth of a nucleated cluster is negligible and approximate the probability density of $t_r$ analytically. The density, $f(t_r)$, has a shape symmetric about odd multiples of the half-period $t_{1/2}$:

$$f(t_r) = g(\Theta) \times \left\{ \begin{array}{ll}
\sinh[(t_r/\langle \tau \rangle) - 2(n-1)\Theta] & \text{for } 2(n-1)t_{1/2} < t_r < (2n-1)t_{1/2} \\
\sinh[2n\Theta - (t_r/\langle \tau \rangle)] & \text{for } (2n-1)t_{1/2} < t_r < 2nt_{1/2}
\end{array} \right. \quad (4)$$

where $\Theta = t_{1/2}/\langle \tau \rangle$, $g(\Theta) = (\langle \tau \rangle)^{-1} e^{-a\Theta}/(1 - e^{-a\Theta})$, and $n = 1, 2, \ldots$. Figure 10 plots observed $f(t_r)$ for $t_{1/2} = 1.2 \times 10^4$. Observed frequencies peak at odd multiples of the half-period, the signature of stochastic resonance (Gammaitoni et al., 1995; Sides et al., 1998b; Acharyya, 1999; Korniss et al., 2002), and match predicted $f(t_r)$ reasonably well.

**DISCUSSION**

Our simulation study examines spatial competition in light of an interaction between the timescale of environmental variation, relative to mean invasion time, and the spatial-clustering pattern underlying those invasions that succeed. When the environment varies slowly relative to the population-dynamic timescale, multi-cluster and single-cluster invasion processes predict the same qualitative behaviour. Each half-period the currently advantaged competitor invades and excludes the resident; the environment never interrupts the endogenous relaxation to equilibrium. However, the timescale-difference interacts with invasion regime when the environmental half-period matches, or is less than, mean invasion
time. We observe a range of dynamics, from exclusion of the rare species to unpredictable exclusion of the resident.

Figure 11A shows time-dependent densities for a rapidly alternating environment in the MC regime. The two species are effectively equivalent competitively, and we observe non-equilibrium co-existence. Decreasing the pace of environmental change to match invasion time (Fig. 11B) generates invasion–exclusion cycles, since the longer half-period often permits relaxation to the currently favoured single-species equilibrium. For the SC regime, Fig. 11C shows the results for rapid environmental variation; the rare species cannot generate a cluster large enough to invade. Finally, increasing the half-period length to match invasion time in the SC regime generates stochastic resonance (see Fig. 11D).

When the environmental and endogenous timescales match, the long-term dynamics is least predictable. In the MC regime, the short-term dynamics is dominated by invasion–exclusion cycles driven by introduction events and alternating competitive advantage. However, randomly (though rarely) the species quickly exchange roles of common and rare. The period-averaged density difference ($Q$) exhibits little short-term variation, but varies bimodally over the long term (Fig. 3E). Hence the norm of the dynamic order parameter $\langle |Q| \rangle$ (see subsection on Metrics) best captures the symmetry breaking associated with the dynamic phase transition from numerical dominance by one species to dominance by its competitor. Deterministic environmental periodicity and stochastic spatial propagation combine to render long-term prediction uncertain for MC invasion. For the SC regime, uncertainty inherent to stochastic resonance is summarized by the distribution of waiting times between switches in roles of common and rare species (Fig. 10).

We began simulations with one species common and the other at zero density, focusing on the dynamics of rarity. We wondered if the invader’s consistent failure to advance in the SC
Fig. 11. Time-dependent densities. $L = 128$. Initial densities $\rho_1(t), \rho_2(t) = (\rho_1^*, 0)$. (A) MC regime; $t_{1/2} = \frac{100}{\langle \tau \rangle}$, non-equilibrium co-existence. (B) MC regime; $t_{1/2} = \frac{800}{\langle \tau \rangle}$, invasion–exclusion cycles. (C) SC regime; $t_{1/2} = \frac{400}{\langle \tau \rangle}$, no invasion observed. (D) SC regime; $t_{1/2} = \frac{8000}{\langle \tau \rangle}$, stochastic resonance.
regime ($t_{1/2} \leq \langle \tau \rangle$) and/or the exclusion–invasion cycles in the MC regime ($t_{1/2} \leq \langle \tau \rangle$) were consequences of initial conditions. We repeated these simulations (only once each) with initial condition $[\rho_1 = 0.5, \rho_2 = 0.48]$. The results, presented in Fig. 12, closely resemble those in Fig. 11, after transients disappear. We conclude that the dynamics does not depend on initial conditions.

Any approximation to our model that assumes strong spatial mixing will fail to produce the observed range of dynamics (Antonovics et al., 2006). Our model’s properties of discreteness and stochasticity define its fundamental character (O’Malley et al., 2009), and the dynamics of rarity is studied realistically by assuming discrete individuals and stochastic demographic events (Durrett and Levin, 1994; Ellner et al., 1998; Duryea et al., 1999; Escudero et al., 2004).

We described our work as examining interaction of the difference between timescales with invasion regime. This categorization organized our simulation study. However, scaling arguments of nucleation theory allow us to simplify the description as interaction between the exogenous timescale $t_{1/2}$ and the introduction rate $\beta$, since both the characteristic timescale $\langle \tau \rangle$ and the characteristic length scale $D$ of the endogenous dynamics depend on $\beta$.

Consider MC invasion in a constant environment, and let habitat size $L$ grow large. Then $\langle \tau \rangle \sim (Iv^2)^{-1/3}$, where $I$ is the nucleation rate/unit area and $v$ is the radial velocity at which nucleated clusters of the superior competitor grow. We noted above that $\langle \tau \rangle \sim \beta^{-1/3}$, since $I \sim \beta$ (see Appendix 1). That is, the characteristic timescale for MC invasion increases as the inverse of the cube root of the introduction rate (Korniss and Caraco, 2005). The characteristic length scale $D$, the expected distance between nucleated invader clusters, scales as $D \sim (vI)^{1/3}$, and we have $D \sim \beta^{-1/3}$ (O’Malley et al., 2006a). That is, the characteristic length scale for MC invasion also increases as the inverse of the cube root of the introduction rate.
Propagation and mortality rates are fixed in a constant environment. From the above, for any fixed habitat size $L$, there is a critical introduction rate $\beta_c$ such that for $\beta < \beta_c$, we have $D \sim \beta^{-1/3}/L$. Then, constant-environment invasion crosses over to the SC regime where $\langle \tau \rangle \sim \beta^{-1}$. Therefore, both the endogenous timescale and invasion mode (given habitat size) depend on the introduction rate, and our results reveal an interaction between $t_{1/2}$ and $\beta$.

Model context

Our model assumes rare, but repeating, introduction of individuals, an assumption prompted by observed invasions of non-native species (Veltman et al., 1996; Loreau and Mouquet, 1999; Sax and Brown, 2000). Not every introduction succeeds; those that do initiate spatially clustered growth. The assumption of spatially explicit introduction contrasts with models treating a rare species as spread uniformly across an environment (Snyder and Chesson, 2004). The period-averaged competitive symmetry of our model’s two species is less realistic. However, the results offer a comparison for spatial competition where a superior species’ advantage varies temporally.

A number of previous models examine competition under temporal variation; this study and that of Chan et al. (2009) differ by taking an individual-based perspective. Chesson’s (1990) study represents a standard class of models for competition under temporal variation. Seed-bank densities of two homogeneously mixing plant species are tracked in discrete time. A varying environment and density dependence affect each species’ annual growth rate/unit density. That is, the fraction of a species’ seeds germinating varies randomly between generations, and competition reduces the seed yield per adult. Analysis focuses on growth-rate responses to competition in good versus bad environments, and on differences in the growth rates’ response to increasing competition across environments. This type of model permits greater between-species differences than do our assumptions; co-existence mechanisms are therefore more general. Model construction assumes a single scale for temporal variation; the environment changes randomly, independently each discrete generation. However, autocorrelation can be used to manipulate the exogenous timescale (Chesson, 1990).

Another approach considers timescale interaction explicitly, but retains the homogeneous-mixing assumption. Abrams (2004) develops a continuous-time model where two consumer species compete for a common resource that has its own dynamics with a periodic growth rate. Trade-offs can promote persistence. If the species with lower efficiency at high resource density has the greater efficiency at low resource density, co-existence can result. Abrams (2004) focuses on the rapidity of the consumers’ demographic responses to temporal variation in resource density, and concludes that interaction of endogenous population dynamics and periodic environmental variation impacts patterns in competitor abundances. Recer et al. (1987) and Cross et al. (2005) make similar points for simpler systems.

Schoolmaster and Snyder (2007; see Snyder, 2007) combine temporal and spatial variation in a model for competing perennial plants occupying a one-dimensional environment. An adult’s seed dispersal and competitive suppression of seedling establishment both decline with distance from the adult. Establishment further depends on an environmental quality that varies periodically both in time and along the linear environment. Each periodicity has its own scale. The authors ask how these fluctuations, compared with environmental constancy, affect the growth rate of a rare species dispersed uniformly along the environment. The results suggest that the impact of environmental variation depends on
its interaction with the competitors’ life-history traits (Schoolmaster and Snyder, 2007), hence interaction with the population dynamic timescale. A series of different models addressing interspecific competition agree that quantitative and qualitative predictions can depend on the interaction of endogenous and exogenous timescales. Chan et al. (2009) model individual-based competition between species. Each species’ birth rate varies seasonally, the species’ respective death rates are fixed, and introduction is not considered explicitly. The authors seek analytical conditions for co-existence. To do so, they relax spatial structuring of offspring dispersal, a key assumption of our study. Given a sufficiently large dispersal neighbourhood, Chan et al. (2009) prove that temporal variation may promote co-existence, which is not found absent the variation. Sufficient conditions for co-existence match those of the homogeneous mean-field approximation to their model.

**Cross-disciplinary integration**

Our results demonstrate that the impact of a periodic environment on spatial competition can depend on the difference between exogenous and endogenous timescales, and that the resulting dynamics can be complex. Similar analyses of interacting timescales have successfully advanced understanding of spatially structured physical systems (Chakrabarti and Acharyya, 1999). Particular examples include time-dependent properties of ferromagnetic thin films in an oscillating external magnetic field (Korniss et al., 2000; Robb et al., 2007, 2008; Buendia and Rikvold, 2008), and behaviour of a catalytic reaction subject to periodic variation in CO pressure (Machado et al., 2005). These systems can exhibit the symmetry breaking we found when the endogenous and exogenous timescales matched in the MC invasion regime. Details of local interactions differ substantially among these physical and ecological models. But they share important dynamical behaviours emerging from timescale interaction. We hope that the convergent predictions suggest a more general understanding.

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Spatial competition under temporal variation


Summarizing model transitions at an arbitrary site $x$ in a constant environment, we have:

$$
0 \rightarrow 1, \quad 0 \rightarrow 2, \quad 1 \rightarrow 0, \quad 2 \rightarrow 0,
$$

(A1)

where 0, 1, 2 indicates whether the site is empty, resident-occupied or invader-occupied, respectively. In a constant environment, simulations always reveal strongly clustered growth of the competitively superior invader. Once an invader cluster reaches the critical radius, that cluster, on average, grows approximately deterministically with radial velocity $v$.

For small environments or for sufficiently small introduction rates, so that $L \ll D$, invasion almost always occurs through spread of a single invader cluster. That is, a habitat size much smaller than the mean distance between clusters (the mean we would observe in an infinite environment) implies the SC regime. In simulation, we confirmed that nucleation of a successful invading cluster is a Poisson process with nucleation rate per unit area $I$. In the SC regime, the lifetime of the resident species $\tau$ is dominated by the lengthy waiting time until the first successful invader cluster nucleates. Differences between competitors’ propagation and mortality rates govern the time required for the cluster to grow and exclude
the nucleation rate per unit area is \( \langle t \rangle = \langle t_n \rangle = (L^2 I)^{-1} \) in the SC regime. The \( \beta \)-dependence of the nucleation rate per unit area is \( I - \beta \). Therefore, \( \langle t_n \rangle = (L^2 \beta)^{-1} \); waiting time for invasion varies inversely with habitat area and the introduction rate.

For SC invasion, the cumulative distribution of invasion times \( \Pr[\tau > t] \), i.e. the probability that the resident’s global density has not decayed to \( \rho^*_n/2 \) by time \( t \) is a modification of the 0-term of a Poisson probability function:

\[
P_{\text{mot}}(t) = \begin{cases} 
1 & \text{for } t \leq t_g \\
\exp[-(t - t_g)/(t_n)] & \text{for } t > t_g
\end{cases}
\] (A2)

In the SC regime, \( t_g - L/v \) is the approximately deterministic growth time until the invading species drives the resident to half its initial density. For very small nucleation rates per unit area, \( \langle t_n \rangle \gg t_g \). Therefore, the lifetime of the resident \( \tau \) is governed by the large average waiting time until the first successful invader cluster nucleates, so that \( \langle \tau \rangle = \langle t_n \rangle + t_g = \langle t_n \rangle = (L^2 I)^{-1} \) in the SC regime.

In large environments (or for large introduction rates), \( \Pr[\tau > t] \) approaches a step-function centred on the system size-independent lifetime \( \langle \tau \rangle \). In the limit of a large environment, we can approximate global densities closely by Avrami’s law, or KJMA theory (Kolmogorov, 1937; Johnson and Mehl, 1939; Avrami, 1940):

\[
\rho_1(t) = \rho^*_n \exp\left[-\ln(2)\left(\frac{t}{\langle \tau \rangle}\right)^{3}\right] \quad \text{and} \quad \rho_2(t) = \rho^*_n \left(1 - \exp\left[-\ln(2)\left(\frac{t}{\langle \tau \rangle}\right)^{3}\right]\right), \quad (A3)
\]

where \( \langle \tau \rangle = (Iv^3)^{-1/3} \). An important result in nucleation theory, Avrami’s law provides the generic functional form of the time-dependent global densities during MC invasion. Furthermore, the parameters of a specific model for spatially structured ecological interactions (i.e. the local transition rates \( \alpha_1, \alpha_2, \mu, \) and \( \beta \)) govern the characteristic timescale (the lifetime \( \langle \tau \rangle \)) through their impact on the nucleation rate per unit area \( I(\alpha_1, \alpha_2, \mu, \beta) \) and the invader-cluster radial velocity \( v(\alpha_1, \alpha_2, \mu) \). Thus, equation (A3) also identifies an important connection between model-specific processes at the level of individual propagation/mortality rates and pattern at the population-dynamic level. We previously linked these two organizational levels through the nucleation rate \( I(\alpha_1, \alpha_2, \mu, \beta) \) and the invader-cluster radial velocity \( v(\alpha_1, \alpha_2, \mu) \) (see O'Malley et al., 2006a, 2009).

Here, we examine a specific model, but the framework of nucleation theory has broad ecological significance. Any time-homogeneous invasion processes combining rare introduction, preemptive competition, and localized propagule dispersal (i.e. strong dispersal limitation) in a large environment will likely generate spatially clustered growth and global dynamics consistent with the predictions of nucleation theory (O’Malley et al., 2006a, 2009; Allstadt et al., 2007). That is, in a constant environment, nucleation theory should predict not only equilibrium states, but also time-dependent global densities, when locally dispersing species compete for space.